

Original research

Computational modelling suggests that Barrett's oesophagus may be the precursor of all oesophageal adenocarcinomas

Kit Curtius (1,2 Joel H Rubenstein,^{3,4} Amitabh Chak,⁵ John M Inadomi⁶

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For numbered affiliations see end of article.

Correspondence to

Kit Curtius, Division of **Biomedical Informatics**, Department of Medicine, University of California, San Diego, La Jolla, CA, USA; kcurtius@health.ucsd.edu

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ABSTRACT

Objective Barrett's oesophagus (BE) is a known precursor to oesophageal adenocarcinoma (OAC) but current clinical data have not been consolidated to address whether BE is the origin of all incident OAC. which would reinforce evidence for BE screening efforts. We aimed to answer whether all expected prevalent BE, diagnosed and undiagnosed, could account for all incident OACs in the US cancer registry data. Design We used a multiscale computational model of OAC that includes the evolutionary process from normal oesophagus through BE in individuals from the US population. The model was previously calibrated to fit Surveillance, Epidemiology and End Results cancer incidence curves. Here, we also utilised age-specific and sex-specific US census data for numbers at-risk. The primary outcome for model validation was the expected number of OAC cases for a given calendar year. Secondary outcomes included the comparisons of resulting model-predicted prevalence of BE and BE-to-OAC progression to the observed prevalence and progression rates.

Results The model estimated the total number of OAC cases from BE in 2010 was 9970 (95% CI: 9140 to 11 980), which recapitulates nearly all OAC cases from population data. The model simultaneously predicted 8%–9% BE prevalence in high-risk males age 45–55, and 0.1%–0.2% non-dysplastic BE-to-OAC annual progression in males, consistent with clinical studies. **Conclusion** There are likely few additional OAC cases arising in the US population outside those expected from individuals with BE. Effective screening of highrisk patients could capture the majority of population destined for OAC progression and potentially decrease mortality through early detection and curative removal of small (pre)cancers during surveillance.



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INTRODUCTION

diagnosed when a patient presents with symptoms such as dysphagia. Unfortunately, the majority of these patients do not live past the first year of their diagnosis because by the time dysphagia develops, metastatic cancer is already present. In order to prevent this cancer or detect it at an earlier, more treatable stage, efforts are now made to identify patients with Barrett's oesophagus (BE), the only known precursor to OAC. Identified BE patients are believed to have a 40-50-fold higher annual

Oesophageal adenocarcinoma (OAC) is typically

Significance of this study

What is already known on this subject?

- Barrett's oesophagus (BE) patients have a 40– 50-fold higher risk of developing oesophageal adenocarcinoma (OAC) than the general population yet many remain undiagnosed.
- Identified BE patients receiving surveillance can have early cancers discovered endoscopically, which decreases the high overall OACassociated mortality.
- Currently, around 90% of patients who develop OAC were never part of a BE surveillance programme, and those BE patients on surveillance have a low annual progression rate of 0.1%-0.3% to develop OAC.

What are the new findings?

- ► By applying a model that incorporates the evolution from normal cells to BE to OAC in patients, we found that the numbers add up-the expected number of OAC cases in the US population are explained by the published rates of BE described above.
- ► We cohesively examined the published estimates to determine that all OAC likely arises from both identified BE and occult, undiagnosed BE in the population.

How might it impact on clinical practice in the foreseeable future?

- Based on current best estimates, our findings suggest that there is no public health need to seek cases of a non-BE alternative pathway to OAC.
- ► Increasing efforts for effective, sensitive screening and surveillance of the true BE population has the potential to decrease OAC mortality in the coming years.

incidence of OAC than the general population.¹ Metaplastic BE progresses through dysplasia to cancer. Advances in endoscopic eradication therapy for dysplastic BE discovered during surveillance of BE can now prevent cancer.² However, most cancers arise in patients without previously diagnosed BE suggesting either inadequate screening strategies or, as a recent study proposes, the possible existence of a pathway independent from the BE pathway.³ In this study, we seek to answer a simple



question about the unseen origins of OAC: does overall OAC incidence reflect the number of cancers that would be expected to arise only from prevalent BE? In other words, do any OAC cases remain unaccounted for that ergo did not arise from the typical Barrett's precursor pathway? The answer to this question will importantly guide research and public health efforts. If BE is the major or only precursor of OAC, then investigators should continue to focus on improving BE detection. If BE is not the major precursor of OAC, then research needs to focus on identifying alternative pathways and BE screening programmes will have limited impact on prevention and early detection of OAC.

In reality, very few individuals who have BE are ever offered an upper endoscopy, and therefore most BE remains asymptomatic and undiagnosed.¹ Patients with gastro-oesophageal reflux disease (GERD) are technically the only subpopulation of the general public typically recommended BE screening because it is believed they have a 5-fold relative risk (RR) of developing long segment BE,⁴ yet even so only about 10% of GERD patients will receive an endoscopy.¹ This indicates underscreening, likely because patients either do not complain of their GERD symptoms, they respond adequately to medical therapy, or were otherwise not deemed suitably high-risk by their physician to warrant an esophagogastroduodenoscopy. Nonetheless, the prevalence of BE in the general population is 1%-2%, whether diagnosed or not,⁵⁶ and this is likely considerably higher in certain at-risk groups in the USA.⁷⁻¹⁰ The main concern is that the average rate to develop OAC in these patients is low-around 0.3% per year.¹¹ Therefore, the majority of endoscopies are futile in finding OAC. We aimed to answer whether all prevalent BE expected, diagnosed and undiagnosed in the US population, could account for all the incident OACs expected as progression rates would imply, to fit the national cancer registry data.

METHODS

The question above is too complex to answer on the 'back of an envelope' because published *average* rates of progression are dependent on age, birth cohort and calendar year. In particular for OAC, age-specific incidence rates vary drastically between men and women.¹² This complexity of timescales involved in normal to premalignant BE to OAC progression has necessitated the creation of quantitative models that analyse cancer incidence rates, and project these trends into the future for public health risk assessments and planning.¹³ Models also quantify the potential impact of progression rates measured in clinical studies on hypothetical intervention and surveillance scheduling in efficacy and cost-effectiveness studies.^{14–16} Such models allow us to perform quantitative, comparative analyses on the benefits versus harms of proposed screening and surveillance protocols against watch-and-wait strategies; these simply cannot be done heuristically due to the complex nature of cancer evolution.

In this study, we model both the onset of BE and the progression of BE to OAC. As a brief background, the multistage clonal expansion model for OAC (herein referred to as the MSCE-OAC model, but also referred to as the MSCE-EAC model elsewhere) is a stochastic model for development of OAC during patient lifetime that includes probabilities of developing BE at various ages, followed by initiation of dysplastic and malignant cell clones in BE with parameters for growth and progression of individual clones to cancer (figure 1). The *inputs* only include GERD prevalence (calibrated to age-specific and sexspecific estimates)^{17 18} and OAC age-specific and sex-specific incidence curves provided by Surveillance, Epidemiology and End Results (SEER) registry.¹² The BE prevalence and neoplastic progression rates are calibrated to fit those inputs, that is, they are not based on observed BE prevalence nor neoplastic progression rates from empiric studies. Briefly, the model includes a GERD-stratified risk curve to develop BE, which is modelled as an age-dependent rate of exponential BE onset each calendar year with an unknown baseline parameter v_0 . The patient-specific BE lengths can vary, derived from a Beta distribution with general population mean length set to 2-3 cm. Beyond v_0 , the baseline constant rate for BE onset, the additional model parameters govern the evolutionary dynamics for dysplastic and malignant growth and OAC detection. The model parameters have been previously calibrated such that the resulting hazard functions fit to OAC age-specific and sex-specific incidence curves provided by SEER registry.¹³ We found during rigorous model selection with likelihood



Figure 1 The stochastic, multiscale model for OAC development (MSCE-OAC) includes conversion from normal squamous epithelium in the oesophagus to BE metaplasia with BE onset rate v(t), which is a function of a baseline rate v₀ and age-dependent prevalence of GERD p_{GERD}(t) (see Methods for details). Two-hit processes with rates μ_0 , μ_1 can initiate a premalignancy (eg, inactivation of tumour suppressor gene *TP53* in non-dysplastic BE due to mutation/copy number alteration in a BE daughter cell creates first cell of a high grade dysplasia lesion). Premalignant cell growth rates are defined as α_p = division rate, β_p = death/differentiation rate per year. Malignant transformation with rate μ_2 creates the first cell of a preclinical clone that can grow with rates α_M = division rate, β_M = death/differentiation rate per year. Size-based probability ρ for detection of preclinical malignant clone can lead to patient-specific time of incident OAC. BE, Barrett's oesophagus; OAC, oesophageal adenocarcinoma; GERD, gastro-oesophageal reflux disease; MSCE-OAC, multistage clonal expansion for oesophageal adenocarcinoma.



Figure 2 The MSCE-OAC model was previously calibrated to SEER incidence curve data stratified by sex and 10-year grouped birth cohorts from 1900 - 1909 to 1950–1959.¹³ The model hazard fits by birth cohort (denoted by colour) represent OAC incidence curves (solid lines) that are consistent with Surveillance, Epidemiology and End Results (SEER) data trends by birth cohort (dashed lines), separately for men (left panel) and women (right panel). MSCE-OAC, multistage clonal expansion for oesophageal adenocarcinoma.

ratio tests that models stratified by birth cohort and sex best fit the incidence data, robust to sensitivity analyses (figure 2). With these fits, the model *outputs* used for this study include the expected number of OAC cases in an at-risk population at a given year calculated using the hazard function h_{OAC} (see online supplemental material for equation details), along with the BE prevalence and the resulting BE-to-OAC progression rates (predicted as specific to age, sex and birth cohort).

This model has been used and improved in comparative analyses within the NCI Cancer Intervention and Surveillance Modelling Network consortium for the past 9 years, which has enabled numerous studies on sensitivity of biopsy sampling techniques for detection of small dysplastic lesions,¹⁴ on influence of patient-specific molecular BE dwell time on future OAC risk,¹⁹ and on cost-effectiveness of endoscopic eradication therapy for certain BE risk groups during surveillance.¹⁵ In our original study on modelling OAC incidence and mortality rates from 1975 to 2010, we used SEER-specific model fits combined with US census data to estimate past and predict future OAC-related deaths but did not include predicted OAC cases by calendar year when applied to US census data.¹³

In the Results below, we expand on prior modelling to help elucidate an answer to our general public health question-'Is BE the precursor of all OAC?' To do this, we first applied the model to estimate the number of OAC cases using the US agespecific and sex-specific at-risk population estimates from the US census data, to be able to compare with the expected number quoted by Vaughan and Fitzgerald.¹ This outcome serves as an independent validation of successful calibration of our model to OAC incidence. Then, we compared the simultaneous predictions of age-specific BE prevalence using the MSCE-OAC model with the published data currently used for screening rationale,²⁰ which included endoscopic reports from the Clinical Outcomes Research Initiative (CORI) for more than 150000 patients, most of whom were born around 1950. We also compared the mathematical predictions of neoplastic progression rate from nondysplastic BE to published estimates.

RESULTS

First, Vaughan and Fitzgerald estimated that the newly diagnosed number of cases for ages greater than 40 to be roughly around 10 000 total in the USA every year based on data from 2010 with an average OAC incidence rate across all age groups.¹ With the Markov model framework, we can analytically compute the OAC hazard function and estimate the expected number of newly diagnosed OAC cases by age and year separately for men and women when considering also population data. As a starting point using 2010 census person-year data,²¹ the model predicts that about 2.2 million adults had prevalent BE in 2010, which is around 1.6% of the general US population over age 40. Then, for age groups greater than 40 in both sexes of all races, our single-age calibrated model estimated that the expected number of new OAC cases diagnosed in 2010 was equal to 9970 (95% CI: 9140 to 11 980).

We also computed the analogous estimate for OAC cases using incidence rates quoted directly from the SEER registry for ages 40–90, which was found to be 9400 OAC cases total in 2010.¹² Thus, the estimate generated by our computational model of progression from BE to OAC is closely consistent with the total number of OAC cases reported in SEER, which also aligns with the 10K incident cases quoted by Vaughn and Fitzgerald.¹ The model therefore suggests that over 90% of OAC cases are attributable to BE.

Second, we considered what the model simultaneously predicted for BE prevalence and BE-to-OAC progression rates in order to achieve the expected ~10K cases. Breaking down the contributions of the 2.2 million total BE patients estimated above, the model predicted BE prevalence to be 1.9%–2.4% in men and 0.4%–0.5% in women in the general US population ages 45–55 in 2010 (figure 3A). These predictions concur with best estimates^{5 6} and influence the total OAC cases predicted by the multistage model. To further explore implications for high-risk patients, we note that the model predicted a BE prevalence of 7.9%–9.3% in US men with symptomatic GERD who are cancer-free ages 45–55 in 2010 when the RR of BE vs non-GERD individuals is assumed to be RR=5 (figure 3B). This is



Figure 3 Model predictions for BE-positive yield in a cancer-free population (solid lines) are consistent with observed data (dashed lines) from Clinical Outcomes Research Initiative (CORI).²⁰ (A) Solid lines show model results for the general US population stratified by sex from the 1950 birth cohort, with contributions of relative risk (RR) of BE from the age-specific, prevalent GERD population assumed to be RR=5 (shaded areas, RR=[2,6]). Dashed lines show consistency with observed BE prevalence data for patients without indication for screening in CORI, which are independent of the model. Model BE prevalence estimates are part of the evolutionary multistage process and thus affect predictions of the total OAC cases predicted (see Results). (B) Solid lines show model results for the symptomatic GERD subpopulation stratified by sex from the 1950 birth cohort with RR for BE set to RR=5. The shaded areas are predicted ranges for GERD and BE length. The true GERD-specific BE prevalence contributing to mathematical formulation used in (A) is within this region, where individual contributions are based on GERD onset age and underlying distribution of RR. Dashed lines show BE prevalence data for patients with GERD, and/or another indication for screening, in CORI. BE, Barrett's oesophagus; GERD, gastrooesophageal reflux disease; OAC, oesophageal adenocarcinoma.

also consistent with the estimate of 8% provided by Vaughan and Fitzgerald¹ for prevalence of cancer-free BE diagnoses among GERD patients who undergo an upper endoscopy. Further, the model's predicted age-specific BE prevalence curves by sex were consistent with previous results on BE prevalence from the CORI study²⁰ (figure 3A). Compared with our model results and 8% quoted above¹ for high-risk groups, the CORI study independently found similar BE prevalence in white men with GERD of 6.3% for ages 40-49 and 9.3% for ages 50-59 (figure 3B). To account for likely heterogenous RR of developing BE in GERD populations based on symptom onset age, BE length and other factors,^{4 22-26} we also considered a range of fixed values (RR=2-6) and found age-specific trends broadly consistent to overall BE prevalence results in CORI. Observed BE prevalence in white women undergoing screening was less precise in the CORI study data yet still coincided with our predictions for women (figure 3B).

In a sensitivity analysis, we also found these results to be robust to varying GERD prevalence in the model input for men and women in the population (see online supplemental material and online supplemental figure S1). When assuming smaller values of RR that lead to reduced BE prevalence in the GERD subpopulations for both sexes (see online supplemental material for details, online supplemental figure S2), the model still predicts that the majority of expected OACs (over 90%) develop in BE patients.

Finally, we previously found using this model that, for individuals born after 1940, the range of progression rates from BE-to-OAC was 0.10%–0.20% for men, and this was about twice as high as we found for women.¹³ These are plausibly low rates compared with current best estimates.^{11 27} Taken together, these secondary outcomes support the plausibility of our model's predictions for numbers of OAC cases from BE annually.

The modelling results above imply that, even in the most conservative probability estimates, less than 10% of all annual OAC cases are unaccounted for beyond those expected to arise from BE. If there were a more significant alternate non-BE pathway than these numbers imply, then this model (which does not include a non-BE pathway) would have estimated either a much lower predicted population incidence of OAC than what was observed in SEER or shown greater inconsistencies with BE studies. In the latter case, the model would have estimated a greater prevalence of BE than what has been observed, and/or a greater rate of neoplastic progression among non-dysplastic BE than observed.

DISCUSSION

Based on the published epidemiology of BE and OAC, our analysis suggests that a major alternative non-BE pathway to OAC is an unlikely scenario. The existence of such an alternative pathway was suggested by a retrospective analysis of macroscopic reports of OAC specimens diagnosed without BE in two cohorts from the USA and UK by Sawas and colleagues; however, their study conclusions remain speculative due to some important limitations including (1) a lack of longitudinally followed cases to OAC from non-BE patient oesophageal tissue and (2) the plausibility that small BE segments were completely overtaken by malignant expansions and thus were unmeasurable at cancer diagnosis.³ Moreover, our result that BE is the main origin of OAC does not necessarily refute the existence of differing phenotypes for OAC-the finding that the presence of BE was associated with better survival could plausibly be explained by the theory that more aggressive cancers are likely to replace the precursor BE more readily than less aggressive cancers. The stochastic nature of our model allows for variation in progression across a population and we explored a wide range of parameter values for rates defining the stochastic process from birth to clinical OAC and reached similar results, but there is still ultimately some uncertainty.

Indeed, genetic and epigenetic analyses have also consistently shown BE and OAC to be very similar,²⁸⁻³¹ and one study that sought genomic differences between adenocarcinomas with and without BE failed to reveal molecular differences between the two.³² Nonetheless, this is fortunate news that, with adequate uptake, screening for BE by upper endoscopy or minimally invasive non-endoscopic technologies^{16 33} could potentially identify and enrol all patients who are at risk for developing OAC into a surveillance programme.

Although the overall progression to OAC is low in patients diagnosed with BE, for those selected BE patients who have high grade dysplasia and/or early OAC detected during surveillance, effective treatment can save lives. In this way, our analysis reinforces the primary goal in BE screening for OAC prevention-that effective surveillance of the entire BE population could potentially prevent the majority of mortality caused by OAC in the general population. Further, by mathematically analysing the time-dependent nature of cumulative risk of BE in GERD patients, we can also use our multistage model framework to improve identification of at-risk populations by optimising the timing of initial screening recommended for BE in symptomatic GERD.³⁴ Although current intensive 'one-sizefits-all' surveillance strategies³⁵⁻⁴⁰ would lead to high costs for those over-diagnosed BE screen cases and surveillance strategies clearly need to improve, we conclude that there is a strong rationale for screening for BE to reduce OAC mortality.

Author affiliations

¹Centre for Genomics and Computational Biology, Barts Cancer Institute, School of Medicine and Dentistry, Queen Mary University of London, London, UK ²Division of Biomedical Informatics, Department of Medicine, University of California

San Diego, La Jolla, California, USA ³Division of Gastroenterology, University of Michigan, Ann Arbor, Michigan, USA

⁴Center for Clinical Management Research, Ann Arbor Veterans Affairs Medical Center, Ann Arbor, Michigan, USA

⁵Division of Gastroenterology and Liver Disease, University Hospitals Cleveland Medical Center, Case Western Reserve University School of Medicine, Cleveland, Ohio, USA

⁶Department of Medicine, University of Utah School of Medicine, Salt Lake City, Utah, USA

Twitter Kit Curtius @yosoykit

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Data availability statement All data relevant to the study are included in the article or uploaded as Supplementary material. All data used in our analysis are

publicly available. CORI data can be accessed through application with ethical approval to NIDDK (https://niddkrepository.org/studies/cori/). All equations are provided either in Figures and Supplementary material or were previously published along with model parameters. Code to solve equations was developed in R (V.3.6.1). Computational scripts are available at: github.com/yosoykit/BEtoEAC_Results.

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ORCID iD

Kit Curtius http://orcid.org/0000-0002-2678-0960

REFERENCES

- Vaughan TL, Fitzgerald RC. Precision prevention of oesophageal adenocarcinoma. Nat Rev Gastroenterol Hepatol 2015;12:243–8.
- 2 Wani S, Qumseya B, Sultan S, *et al*. Endoscopic eradication therapy for patients with Barrett's esophagus-associated dysplasia and intramucosal cancer. *Gastrointest Endosc* 2018;87:907–31.
- 3 Sawas T, Killcoyne S, Iyer PG, et al. Identification of prognostic phenotypes of esophageal adenocarcinoma in 2 independent cohorts. *Gastroenterology* 2018;155:1720–8.
- 4 Taylor JB, Rubenstein JH. Meta-Analyses of the effect of symptoms of gastroesophageal reflux on the risk of Barrett's esophagus. *Am J Gastroenterol* 2010;105:1730–7.
- 5 Zagari RM, Fuccio L, Wallander M-A, et al. Gastro-oesophageal reflux symptoms, oesophagitis and Barrett's oesophagus in the general population: the Loiano-Monghidoro study. Gut 2008;57:1354–9.
- 6 Ronkainen J, Aro P, Storskrubb T, et al. Prevalence of Barrett's esophagus in the general population: an endoscopic study. Gastroenterology 2005;129:1825–31.
- 7 Ward EM, Wolfsen HC, Achem SR, et al. Barrett's esophagus is common in older men and women undergoing screening colonoscopy regardless of reflux symptoms. Am J Gastroenterol 2006;101:12–17.
- 8 Rex DK, Cummings OW, Shaw M, et al. Screening for Barrett's esophagus in colonoscopy patients with and without heartburn. *Gastroenterology* 2003;125:1670–7.
- 9 Gerson LB, Shetler K, Triadafilopoulos G. Prevalence of Barrett's esophagus in asymptomatic individuals. *Gastroenterology* 2002;123:461–7.
- Rubenstein JH, Morgenstern H, Appelman H, et al. Prediction of Barrett's esophagus among men. Am J Gastroenterol 2013;108:353–62.
- 11 Kroep S, Lansdorp-Vogelaar I, Rubenstein JH, et al. An accurate cancer incidence in Barrett's esophagus: a best estimate using published data and modeling. Gastroenterology 2015;149:577–85.
- 12 SEER*Explorer: An interactive website for SEER cancer statistics. Surveillance research program, National cancer Institute. Available: https://seer.cancer.gov/explorer/ [Accessed 15 Apr 2019].
- 13 Kong CY, Kroep S, Curtius K, et al. Exploring the recent trend in esophageal adenocarcinoma incidence and mortality using comparative simulation modeling. *Cancer Epidemiol Biomarkers Prev* 2014;23:997–1006.
- 14 Curtius K, Hazelton WD, Jeon J, *et al*. A multiscale model evaluates screening for neoplasia in Barrett's esophagus. *PLoS Comput Biol* 2015;11:e1004272.
- 15 Kroep S, Heberle CR, Curtius K, et al. Radiofrequency ablation of Barrett's esophagus reduces esophageal adenocarcinoma incidence and mortality in a comparative modeling analysis. Clin Gastroenterol Hepatol 2017;15:1471–4.
- 16 Heberle CR, Omidvari A-H, Ali A, et al. Cost effectiveness of screening patients with gastroesophageal reflux disease for Barrett's esophagus with a minimally invasive cell sampling device. *Clin Gastroenterol Hepatol* 2017;15:1397–404.
- 17 Ruigómez A, García Rodríguez LA, Wallander M-A, *et al.* Natural history of gastrooesophageal reflux disease diagnosed in general practice. *Aliment Pharmacol Ther* 2004;20:751–60.
- 18 Ruigómez A, Wallander M-A, Lundborg P, et al. Gastroesophageal reflux disease in children and adolescents in primary care. Scand J Gastroenterol 2010;45:139–46.
- 19 Curtius K, Wong C-J, Hazelton WD, et al. A molecular clock infers heterogeneous tissue age among patients with Barrett's esophagus. *PLoS Comput Biol* 2016;12:e1004919.

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- 20 Rubenstein JH, Mattek N, Eisen G. Age- and sex-specific yield of Barrett's esophagus by endoscopy indication. *Gastrointest Endosc* 2010;71:21–7.
- 21 United States Census Bureau. Census of population and housing, 2015. Available: http://www.census.gov/prod/www/decennial.html [Accessed Feb 2020].
- 22 Thrift AP, Kramer JR, Qureshi Z, et al. Age at onset of GERD symptoms predicts risk of Barrett's esophagus. Am J Gastroenterol 2013;108:915–22.
- 23 Eisen GM, Sandler RS, Murray S, et al. The relationship between gastroesophageal reflux disease and its complications with Barrett's esophagus. Am J Gastroenterol 1997;92:27–31.
- 24 Anderson LA, Watson RGP, Murphy SJ, et al. Risk factors for Barrett's oesophagus and oesophageal adenocarcinoma: results from the FINBAR study. World J Gastroenterol 2007;13:1585–94.
- 25 Edelstein ZR, Bronner MP, Rosen SN, et al. Risk factors for Barrett's esophagus among patients with gastroesophageal reflux disease: a community clinic-based case-control study. Am J Gastroenterol 2009;104:834–42.
- 26 Shaheen N, Ransohoff DF. Gastroesophageal reflux, Barrett esophagus, and esophageal cancer: scientific review. *JAMA* 2002;287:1972–81.
- 27 de Jonge PJF, van Blankenstein M, Looman CWN, et al. Risk of malignant progression in patients with Barrett's oesophagus: a Dutch nationwide cohort study. Gut 2010;59:1030–6.
- 28 Ross-Innes CS, Becq J, Warren A, et al. Whole-genome sequencing provides new insights into the clonal architecture of Barrett's esophagus and esophageal adenocarcinoma. Nat Genet 2015;47:1038–46.
- 29 Secrier M, Li X, de Silva N, et al. Mutational signatures in esophageal adenocarcinoma define etiologically distinct subgroups with therapeutic relevance. Nat Genet 2016;48:1131–41.
- 30 Yu M, Maden SK, Stachler M, *et al.* Subtypes of Barrett's oesophagus and oesophageal adenocarcinoma based on genome-wide methylation analysis. *Gut* 2019;68:389–99.

- 31 Liu Y, Sethi NS, Hinoue T, et al. Comparative molecular analysis of gastrointestinal adenocarcinomas. Cancer Cell 2018;33:721–35.
- 32 Ferrer-Torres D, Nancarrow DJ, Kuick R, et al. Genomic similarity between gastroesophageal junction and esophageal Barrett's adenocarcinomas. Oncotarget 2016;7:54867–82.
- 33 Moinova HR, LaFramboise T, Lutterbaugh JD, et al. Identifying DNA methylation biomarkers for non-endoscopic detection of Barrett's esophagus. Sci Transl Med 2018;10:aao5848.
- 34 Curtius K, Dewanji A, Hazelton WH, et al. Optimal timing for cancer screening and adaptive surveillance using mathematical modeling. *bioRxiv* 2020.
- 35 Codipilly DC, Chandar AK, Singh S, *et al.* The effect of endoscopic surveillance in patients with Barrett's esophagus: a systematic review and meta-analysis. *Gastroenterology* 2018;154:2068–86.
- 36 Spechler SJ, Sharma P, Souza RF, et al. American Gastroenterological Association Technical Review on the Management of Barrett's Esophagus. Gastroenterology 2011;140:e18–52.
- 37 Shaheen NJ, Falk GW, Iyer PG, et al. American College of Gastroenterology. ACG Clinical Guideline: Diagnosis and Management of Barrett's Esophagus. Am J Gastroenterol 2016;111:30–50.
- 38 Fitzgerald RC, di Pietro M, Ragunath K, et al. British Society of Gastroenterology guidelines on the diagnosis and management of Barrett's oesophagus. Gut 2014;63:7–42.
- 39 di Pietro M, Fitzgerald RC, BSG Barrett's guidelines working group. Revised British Society of Gastroenterology recommendation on the diagnosis and management of Barrett's oesophagus with low-grade dysplasia. *Gut* 2018;67:392–3.
- 40 Qumseya B, Sultan S, Yang J, et al. ASGE guideline on screening and surveillance of Barrett's esophagus. *Gastrointest Endosc* 2019;90:335–59.

Supplementary Material

Model hazard function and expected number of esophageal adenocarcinoma (EAC) cases

We compared the model's predictions to the US population estimates for Barrett's esophagus (BE) and EAC quoted by Vaughan and Fitzgerald,¹ by combining age- and sex-specific model hazard rates with at-risk population estimates from US census data.² For expected EAC incidence, using the Markov model framework (see Figure 1) we can analytically compute the EAC hazard function h_{EAC} (see Curtius et al. for full derivation with age-specific GERD-dependent BE rates incorporated explicitly ³) and estimate the expected number of newly diagnosed EAC cases by age and year separately for men and women, using population data for the at-risk population numbers.² This is computed as:

$$\Lambda_{i,j} = PY_{i,j}h_{EAC}(a_i, b_k)$$

where $PY_{i,j}$ is the number of person-years at-risk in period c_j of age a_i and birth cohort $b_k = c_j - a_i$. For the Age-Cohort model, the birth cohort specific hazard h_{EAC} (previously fit using person-year data for specified US populations directly from SEER⁴) can be written as $h_{EAC}(a_i, b_k) = h_{EAC}(t \mid t = a_i, b_k)$. For the main Results of incident EAC cases in $c_j = 2010$ and ages a_i between 40-90, this was computed separately for men and women and the 95% confidence interval for this estimate of summed total cases was computed by re-sampling Markov Chain Monte Carlo posterior distributions of birth year- and sex-specific model parameter estimates^{3, 4} for 100K bootstrap iterations. The equation below calculates the expected total number of EAC cases diagnosed in 2010, Λ_{2010} , which was equal to,

$$\sum_{i=40}^{90} \Lambda_{male_{i,2010}} + \Lambda_{female_{i,2010}} = \mathbf{9}, \mathbf{970} \ [95\% \text{ CI}: \ 9, 140 - 11, 980]$$

For the analogous calculation using SEER incidence rates extracted from SEER*Explorer ⁵ for ages 40-90 in 2010 and census person-year data² we estimated 9,400 EAC cases total.

Sensitivity analysis: model input for gastroesophageal reflux disease (GERD) prevalence

The MSCE-EAC model *inputs* only include age- and sex- specific GERD prevalence and EAC incidence curves provided by SEER registry. Briefly, we previously developed a GERD model with GERD prevalence increasing in accordance with the data for age-specific GERD incidence (see Kong et al. 2014 for more detail ⁴). The model also includes a parameter representing reversion of GERD symptoms, allowing us to fit age-adjusted GERD prevalence based on the US population between ages 40 to 85 to an approximate target of 20%, consistent with population-based studies of GERD prevalence. ⁶ This baseline input is provided in Figure S1A,B (solid blue line for males, and solid red line for females).

To determine the effect of varying GERD prevalence on the results for BE prevalence, we performed an additional sensitivity analysis to provide results based on explorations for a range of age-specific GERD prevalence, specifically between a 50% increase and decrease of values of the baseline function. Even with this range of GERD prevalence used as input (see Figure S1A,B), we found that the model's predictions for BE prevalence are robust to

changes in GERD prevalence with a narrow range of resulting BE prevalence produced as output for both men and women in the total population, which also still correspond well

with CORI BE prevalence estimates (Figure S1C,D).



Figure S1: Sensitivity analysis for age- and sex- specific GERD prevalence model input in men (A) and women (B). With this input, corresponding estimates of BE prevalence in non-EAC population for main MSCE-EAC model results with total population considered are shown for men (C) and women (D) with RR=5. Shaded regions show general consistency with CORI data, and are similar to baseline Results in the main Text.

Sensitivity analysis: model input for relative risk (RR) of BE in GERD vs. non-GERD populations

We found that the model is very robust to changes in relative risk (RR) of BE among individuals with GERD symptoms compared with individuals without GERD symptoms. In the right panel of Figure S2, prevalence of BE in the GERD symptomatic population is displayed (governed by a GERD-specific BE rate $v_{GERD} = RR^*v_0$ derived from the total model). When we assume RR to be equal to 4, which reduces the BE prevalence among individuals with GERD symptoms to 7.5% at age 55 and 8.7% at age 75 (right panel of Figure S2, grey dashed lines), the total number of EAC cases are estimated to be 9,904 [95% CI 9,037- 11,848]. Moreover, reducing the RR to 3 lowers the BE prevalence of men with GERD symptoms to 5.7% at age 55 and 6.6% at age 75 years (right panel of Figure S2, grey dotted lines) and the total number of EAC cases are estimated to be 9,843 [95% CI 8,930 – 11,719]. The reason the model is so robust in total number of predicted EACs is because the predictions for total population BE prevalence (Figure 3A; see left panel of Figure S2 below) are not very sensitive to changes in RR of GERD for BE (grey lines illustrate the narrow predicted range for BE prevalence), where the total model structure considers the age-dependency of the prevalence of GERD symptoms, $p_{GERD}(t)$, in the time-dependent BE development rate, $v(t) = v_0[p_{GERD}(t)*RR+(1-p_{GERD}(t))]$, with baseline BE rate v_0 .



Figure S2: Sensitivity analysis for model predicted BE prevalence when lowering relative risk (RR) of BE in GERD versus non-GERD from the baseline value RR=5 (solid lines, see Figure 3 main Text) to RR=4 (dashed grey lines) and RR=3 (dotted grey lines). The total model predicted BE prevalence in men and women (left panel) is robust to changes in input for RR that consequently lower corresponding BE prevalence in GERD subpopulations (right panel).

Patient and Public Involvement

We did not directly include PPI in this study, but the SEER database used here is updated by an NCI committee that seeks quality improvement with patient/public feedback welcomed.

References

- 1. Vaughn TL, et al. *Nat Reviews Gastroenterol Hepatol* 2015;12:243-248.
- 2. United States Census Bureau. *Census of Population and Housing* [online], http:// www.census.gov/prod/www/decennial.html (2015). Accessed Feb 2020.
- 3. Curtius K, et al. *PLoS Comput Biology* 2016 May 11;12(5):e1004919.
- 4. Kong CY, et al. *Cancer Epidemiol Biomarkers Prev* 2014;23(6):997-1006.
- 5. SEER*Explorer: An interactive website for SEER cancer statistics [online], https://seer.cancer.gov/explorer/ (2019).
- 6. El-Serag HB, et al. Update on the epidemiology of gastro-oesophageal reflux disease: a systematic review. *Gut* 2014;63(6):871-880.